

1. CASINO EXAMPLE

Assume the transition matrix is

$$A = \begin{bmatrix} I & F & L \\ IF & 0.95 & 0.05 \\ DL & 0.1 & 0.9 \end{bmatrix}$$

and the emission probability matrix is

$$B = \begin{bmatrix} 1 & 2 & 3 & 4 & 5 & 6 \\ F & \frac{1}{6} & \frac{1}{6} & \frac{1}{6} & \frac{1}{6} & \frac{1}{6} \\ L & \frac{1}{10} & \frac{1}{10} & \frac{1}{10} & \frac{1}{10} & \frac{1}{10} \end{bmatrix}$$

in which "F" denotes fair die and "L" represents loaded die. Denote Y as the number of dies, and X the status of the die, i.e., $X = 0$ means loaded and $X = 1$ indicates fair. If we have the observation $Y = \{6, 6, 6\}$, use maximum likelihood and maximum *a posteriori* estimate to estimate the status of the die.

解:

$$\begin{cases} P(X=0) = P(X=0) \cdot 0.9 + P(X=1) \cdot 0.05 \\ P(X=1) = P(X=0) \cdot 0.1 + P(X=1) \cdot 0.95 \end{cases} \Rightarrow \begin{cases} P(X=0) = \frac{1}{3} \\ P(X=1) = \frac{2}{3} \end{cases}$$

① MLE maximize $P(Y=(6,6,6)|X)$

$$\begin{aligned} P(Y=(6,6,6)|X=(0,0,0)) &= \frac{1}{2} \cdot \frac{1}{2} \cdot \frac{1}{2} = \frac{1}{8} & P(Y=(6,6,6)|X=(0,0,1)) &= \frac{1}{2} \cdot \frac{1}{2} \cdot \frac{1}{6} = \frac{1}{24} \\ P(Y=(6,6,6)|X=(0,1,0)) &= \frac{1}{2} \cdot \frac{1}{6} \cdot \frac{1}{2} = \frac{1}{24} & P(Y=(6,6,6)|X=(0,1,1)) &= \frac{1}{2} \cdot \frac{1}{6} \cdot \frac{1}{6} = \frac{1}{72} \\ P(Y=(6,6,6)|X=(1,0,0)) &= \frac{1}{6} \cdot \frac{1}{2} \cdot \frac{1}{2} = \frac{1}{24} & P(Y=(6,6,6)|X=(1,0,1)) &= \frac{1}{6} \cdot \frac{1}{2} \cdot \frac{1}{6} = \frac{1}{72} \\ P(Y=(6,6,6)|X=(1,1,0)) &= \frac{1}{6} \cdot \frac{1}{6} \cdot \frac{1}{2} = \frac{1}{72} & P(Y=(6,6,6)|X=(1,1,1)) &= \frac{1}{6} \cdot \frac{1}{6} \cdot \frac{1}{6} = \frac{1}{216} \end{aligned}$$

$$\text{So } \hat{X}_{\text{ML}} = (0, 0, 0)$$

② MAP maximize $P(X|Y=(6,6,6)) = \frac{P(X, Y=(6,6,6))}{P(Y=(6,6,6))} = \frac{P(Y=(6,6,6)|X) \cdot P(X)}{P(Y=(6,6,6))}$

$$\begin{aligned} P(X=(0,0,0)) &= P(X=0) \cdot 0.9 \cdot 0.9 = 0.27 & \therefore P(X=(0,0,0), Y=(6,6,6)) &= \frac{1}{8} \cdot 0.27 = \frac{27}{800} = 0.03375 \\ P(X=(0,0,1)) &= P(X=0) \cdot 0.9 \cdot 0.1 = 0.03 & \therefore P(X=(0,0,1), Y=(6,6,6)) &= \frac{1}{24} \cdot 0.03 = \frac{1}{800} \\ P(X=(0,1,0)) &= P(X=0) \cdot 0.1 \cdot 0.05 = \frac{1}{600} & \therefore P(X=(0,1,0), Y=(6,6,6)) &= \frac{1}{14400} \cdot \frac{1}{600} = \frac{1}{86400} \\ P(X=(0,1,1)) &= P(X=0) \cdot 0.1 \cdot 0.95 = \frac{19}{600} & \therefore P(X=(0,1,1), Y=(6,6,6)) &= \frac{1}{72} \cdot \frac{19}{600} = \frac{19}{43200} \\ P(X=(1,0,0)) &= P(X=1) \cdot 0.05 \cdot 0.9 = 0.03 & \therefore P(X=(1,0,0), Y=(6,6,6)) &= \frac{1}{24} \cdot 0.03 = \frac{1}{800} \\ P(X=(1,0,1)) &= P(X=1) \cdot 0.05 \cdot 0.1 = \frac{1}{300} & \therefore P(X=(1,0,1), Y=(6,6,6)) &= \frac{1}{72} \cdot \frac{1}{300} = \frac{1}{21600} \\ P(X=(1,1,0)) &= P(X=1) \cdot 0.95 \cdot 0.05 = \frac{19}{600} & \therefore P(X=(1,1,0), Y=(6,6,6)) &= \frac{1}{72} \cdot \frac{19}{600} = \frac{19}{43200} \\ P(X=(1,1,1)) &= P(X=1) \cdot 0.95 \cdot 0.95 = \frac{361}{600} & \therefore P(X=(1,1,1), Y=(6,6,6)) &= \frac{1}{216} \cdot \frac{361}{600} = \frac{361}{129600} \approx 2.79 \times 10^{-3} \end{aligned}$$

$$\text{So } \hat{X}_{\text{MAP}} = (0, 0, 0)$$

2. DIFFERENT ESTIMATES

- (1) Suppose that $z = s + v$, where s and v are independent, jointly distributed RVs with $s \sim \mathcal{N}(\eta, \sigma_s^2)$ and $v \sim \mathcal{N}(0, V^2)$.
 - (a) Derive an expression for $E[s|z=z]$.
 - (b) Derive an expression for $E[s^2|z=z]$.
- (2) Suppose that $z = s + v$, where s and v are independent, jointly distributed RVs with $s \sim \mathcal{N}(\eta_s, \sigma_s^2)$ and $v \sim \mathcal{N}(0, \sigma_v^2)$. Assume we have measurements $z(1), \dots, z(n)$.
 - (a) Derive the maximum likelihood estimate for s ;
 - (b) Derive the maximum *a posteriori* estimate for s ;
 - (c) Derive the minimum mean square estimate for s ;
 - (d) Derive the linear minimum mean square estimate for s ;
 - (e) Derive the least squares estimate for s provided measurements $z(1), \dots, z(n)$;
 - (f) Suppose at each time k ($k \in \{1, \dots, n\}$), there is a new measurement $z(k)$, derive the recursive least squares estimate for s . (Assume $\hat{s}_0 = E(s)$, the initial error covariance is P_0);
 - (g) Compare all these 6 kinds of estimates.

$$(1) (a) E(z) = E(s+v) = \eta$$

$$\text{Var}(z) = \text{Var}(s+v) = \text{Var}(s) + \text{Var}(v) = 6^2 + V^2$$

from b)

$$f_z(z) = \frac{1}{\sqrt{2\pi} \sqrt{6^2 + V^2}} e^{-\frac{1}{2} \frac{(z-\eta)^2}{6^2 + V^2}}$$

$$f(s|z) = \frac{f(s, z)}{f(z)}$$

$$\text{Cov}(s, z) = E(sz) - E(s)E(z)$$

$$E(sz) = E(s^2 + sv) = E(s^2) + E(sv) = \text{Var}(s) + E(s)^2 + \text{Cov}(sv) + E(s)E(v)$$

$$= 6^2 + \eta^2$$

$$\text{Therefore } \text{Cov}(s, z) = 6^2$$

$$\text{Therefore } \Sigma = \begin{bmatrix} 6^2 + \eta^2 & 6^2 \\ 6^2 & 6^2 \end{bmatrix}$$

$$\therefore f(s, z) = \frac{1}{2\pi\sqrt{6V}} \cdot \exp\left\{-\frac{1}{2} \cdot [z-s] \Sigma^{-1} \begin{bmatrix} z \\ s \end{bmatrix}\right\}$$

$$= \frac{1}{2\pi\sqrt{6V}} \exp\left\{-\left[\frac{(z-s)^2}{2V^2} + \frac{(s-\eta)^2}{26^2}\right]\right\}$$

$$\therefore f(s|z) = \frac{f(s, z)}{f(z)} = \frac{1}{\sqrt{2\pi} \sqrt{\frac{6V}{6^2 + V^2}}} \cdot \exp\left\{-\frac{1}{2} \frac{(s - [\eta + \frac{6^2}{6^2 + V^2}(z - \eta)])^2}{6^2 + V^2}\right\}$$

$$\therefore E(s|z) = \eta + \frac{6^2}{6^2 + V^2}(z - \eta)$$

$$\text{Var}(s|z) = \frac{6^2V^2}{6^2 + V^2}$$

$$(b) \text{Var}(s|z) = E(s^2|z) - E(s|z)^2$$

$$\therefore E(s^2|z) = \text{Var}(s|z) + E(s|z)^2$$

$$= \frac{6^2V^2}{6^2 + V^2} + \frac{6^4}{(6^2 + V^2)^2} z^2 + 2 \frac{6^2V^2}{(6^2 + V^2)^2} z\eta + \frac{V^4}{(6^2 + V^2)^2} \eta^2$$

(2)

$$z = s + v \quad E(z) = E(s) + E(v) = \eta_s \quad \text{Var}(z) = E(z) + E(v) = 6s^2 + 6v^2 \quad E(z^2) = \text{Var}(z) + E(z)^2 = 6s^2 + 6v^2 + \eta_s^2$$

$$E(sz) = E(s^2 + sv) = 6s^2 + \eta_s^2 \quad \text{Cov}(sz) = E(sz) - E(s)E(z) = 6s^2$$

$$\text{全 } X = \begin{bmatrix} s \\ z_1 \\ \vdots \\ z_n \end{bmatrix}_{n \times 1} \quad \Sigma = \begin{bmatrix} 6s^2 & 6s^2 & \cdots & 6s^2 \\ 6s^2 & 6s^2 + 6v^2 & & \\ \vdots & & \ddots & \\ 6s^2 & & & 6s^2 + 6v^2 \end{bmatrix}_{n+1} \quad \det \Sigma = (6s^2 + 6v^2)^n \cdot (6s^2 - \frac{n6s^2}{6s^2 + 6v^2})$$

$$u = \begin{bmatrix} \eta_s \\ \vdots \\ \eta_s \end{bmatrix}_{n \times 1}$$

$$\therefore f(z_1, z_2, \dots, z_n, s) = \frac{1}{(\sqrt{2\pi})^{n+1} (\det \Sigma)^{\frac{1}{2}}} \cdot \exp\left\{-\frac{1}{2} \cdot (x-u)^T \Sigma^{-1} (x-u)\right\}$$

$$f(s) = \frac{1}{\sqrt{2\pi} 6s} \exp\left\{-\frac{1}{2} \frac{(s-\eta_s)^2}{6s^2}\right\}$$

$$f(z_1, z_2, \dots, z_n) = \frac{1}{(\sqrt{2\pi})^n (6s^2 + 6v^2)^{\frac{n}{2}}} \exp\left\{-\sum_{i=1}^n \frac{(z_i - \eta_s)^2}{26s^2 + 26v^2}\right\}$$

(a) Maximizes $f(z_1, z_2, \dots, z_n|s)$

$$f(z_1, z_2, \dots, z_n|s) = \frac{f(z_1, z_2, \dots, z_n, s)}{f(s)} = \frac{1}{(\sqrt{2\pi})^n (6s^2 + 6v^2)^{\frac{n}{2}} \sqrt{1 - \frac{n6s^2}{6s^2 + 6v^2}}} \exp\left\{-\frac{1}{2} (x-u)^T \Sigma^{-1} (x-u) + \frac{1}{2} \frac{(s-\eta_s)^2}{6s^2}\right\}$$

$$\frac{\partial \ln f(z_1, z_2, \dots, z_n|s)}{\partial s} = 0 \Rightarrow \frac{2[(x-u)^T \Sigma^{-1} (x-u) - \frac{(s-\eta_s)^2}{6s^2}]}{2s} = 0 \Rightarrow 2(x-u)^T \Sigma^{-1} \begin{bmatrix} 1 \\ 0 \\ \vdots \\ 0 \end{bmatrix} - \frac{2(s-\eta_s)}{6s^2} = 0$$

$$\therefore \hat{s}_{ML} = \frac{1}{n} (z_1 + z_2 + \dots + z_n)$$

(b) Maximizes $f(s|z_1, z_2, \dots, z_n)$

$$f(s|z_1, z_2, \dots, z_n) = \frac{f(z_1, z_2, \dots, z_n, s)}{f(z_1, z_2, \dots, z_n)} \quad \frac{\partial f(s|z_1, z_2, \dots, z_n)}{\partial s} = 0 \Leftrightarrow \frac{\partial \ln f(z_1, z_2, \dots, z_n, s)}{\partial s} = 0$$

$$\Leftrightarrow \frac{2[(x-u)^T \Sigma^{-1} (x-u)]}{2s} = 0$$

$$S = 2(X - \mu)^T \Sigma^{-1} \begin{bmatrix} 1 \\ 0 \end{bmatrix} = 0$$

$$\therefore \hat{s}_{MAP} = \eta_s + \frac{6s^2(z_1 + \dots + z_n) - n\eta_s^2}{6s^2 + 6v^2}$$

(c) $E(s|z_1, z_2, \dots, z_n)$

$$\begin{aligned} f(s|z_1, z_2, \dots, z_n) &= \frac{f(z_1, z_2, \dots, z_n, s)}{f(z_1, z_2, \dots, z_n)} \\ &= \frac{1}{\sqrt{2\pi v^2}} \cdot \exp \left\{ -\frac{1}{2} \frac{[s - \eta_s - \frac{6s^2}{6s^2 + 6v^2}(z_1 + \dots + z_n - n\eta_s)]^2}{6s^2 + 6v^2} \right\} \\ \therefore \hat{s}_{MMSE} &= E(s|z_1, \dots, z_n) = \eta_s + \frac{6s^2}{6s^2 + 6v^2}(z_1 + \dots + z_n - n\eta_s) \end{aligned}$$

(d) $\hat{s} = \lambda z$

$$z = \frac{1}{n}(z_1 + z_2 + \dots + z_n)$$

$$E(sz) = 6s^2 + \eta_s^2 \quad E(z^2) = 6s^2 + 6v^2 + \eta_s^2$$

$$\begin{aligned} \hat{s}_{LMSE} &= \frac{E(sz)}{E(z^2)} \cdot z \\ &= \frac{6s^2 + \eta_s^2}{6s^2 + 6v^2 + \eta_s^2} \cdot \frac{1}{n}(z_1 + z_2 + \dots + z_n) \end{aligned}$$

(e)

$$\hat{s}_{LS} = (H^T H)^{-1} H^T [z_1, \dots, z_n]^T = \frac{1}{n}(z_1 + z_2 + \dots + z_n)$$

$$\begin{cases} K_k = P_{k-1} H_k^T (H_k P_{k-1} H_k^T + R_k)^{-1} \\ \hat{s}_k = \hat{s}_{k-1} + K_k (y_k - H_k \hat{s}_{k-1}) \\ P_k = (I - K_k H_k) P_{k-1} (I - K_k H_k)^T \end{cases}$$

Assume $R_k = R$

$$\begin{aligned} \hat{s}_0 &= E(s) = \eta_s \quad H_k = I \\ K=1 &\quad \left\{ \begin{array}{l} K_1 = P_0 (P_0 + R)^{-1} \\ \hat{s}_1 = \hat{s}_0 + \frac{P_0}{P_0 + R} (z_1 - \hat{s}_0) \\ P_1 = \frac{P_0 R}{P_0 + R} \end{array} \right. \quad \xrightarrow{K=2} \quad \left\{ \begin{array}{l} K_2 = \frac{P_1}{P_1 + R} = \frac{P_0}{2P_0 + R} \\ P_2 = \frac{P_1 R}{P_1 + R} = \frac{P_0 R}{2P_0 + R} \\ \hat{s}_2 = \hat{s}_1 + \frac{P_1}{P_1 + R} (z_2 - \hat{s}_1) \end{array} \right. \quad \dots \\ K=n &\quad \left\{ \begin{array}{l} P_{n-1} = \frac{P_0 R}{(n-1)P_0 + R} \\ K_n = \frac{P_0}{nP_0 + R} \end{array} \right. \quad \hat{s}_n = \hat{s}_{n-1} + K_n (z_n - \hat{s}_{n-1}) \\ &\quad = \frac{(n-1)P_0 R}{nP_0 + R} \hat{s}_{n-1} + \frac{P_0}{nP_0 + R} z_n \end{aligned}$$

(g) MLE: 无需先验信息，在小样本情况下可能不稳定

MAP: 充分考虑先验信息，对于估计问题的不确定性能更好处理

MMSE: 考虑估计误差的平方，适用于线性和非线性

LMSE: 原理基于 MMSE，但仅适用于线性模型

LSE: 寻找一个参数估计，使观测数据与模型的预测值的平方误差之和最小，适用于回归分析和曲线拟合等问题

RLSE: 一种递归估计方法，基于 LSE 的原理，逐步更新参数估计以适应动态系统。